BCN Statistics Course

Repeated measures ANOVA

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BCN Statistics Course Repeated measures ANOVA



Usual ANOVA

2 Repeated measures (M)ANOVA

- Main ideas
- Testing effects in RM-ANOVA: Univariate test
- Testing effects in RM-MANOVA: Multivariate test
- Assumptions of RM-(M)ANOVA
- Use univariate or multivariate?
- Decomposing the within-factor effect
- Post hoc procedures
- Extending RM-ANOVA

'Usual' ANOVA

Generalization of independent samples *t*-test to more than 2 groups.

One-way ANOVA (Stats II)

- One DV Y, one categorical IV A
- Compares the means of Y among all groups defined by A

Two-way ANOVA (Stats II)

- One DV Y, two categorical IVs A, B (factors)
- Compares the means of Y among all groups jointly defined by A, B
- It analyzes the main effect of A (across all levels of B), main effect of B (across all levels of A), and interaction effect A × B ('differences between differences')

'Usual' ANOVA: Assumptions

Assumptions in ANOVA:

- Independence of observations.
- Groups defined by factors are normally distributed.
- Groups defined by factors have equal population variances (homogeneity of variance).

Obs: One-way and two-way ANOVA are just special regression models, where the factors are dummy-coded into the model (Stats II & Stats III).

'Usual' ANOVA: Limitation

- So far, we only allowed each subject to contribute with one score (one measurement) on the DV:
 "Different subjects contribute to different means."
- What if each subject is measured in ALL levels of the treatment factor?

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Repeated measures (M)ANOVA

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- Situation 1 A group of subjects is measured under various treatment conditions (or at different points in time).
- Situation 2 Subjects are grouped according to their gender and measured under various treatment conditions (or at different points in time).
- Situation 3 Two groups of subjects are administered two types of drugs at each of three doses.
- Situation 4 Subjects are grouped according to their age (2 groups) and diet program (3 groups), and their weight loss is measured in three different moments (2,4,6 months after beginning of diet).

Main ideas

Repeated measures (M)ANOVA



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Subjects are now measured more than once (for different treatments/at different points in time): Repeated measures.
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- Between-factors: Variables grouping subjects.
- Within-factors: Repeated measures variables.

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- Between-factors: Variables grouping subjects.
- Within-factors: Repeated measures variables.

Situation	Between factor(s)	Within factor(s)	Design 💌 —
Situation 1	—	Treatments	one within
Situation 2	Gender	Treatments	one between, one within
Situation 3	Group	Drug, Dosage	one between, two within
Situation 4	Age, Diet program	Weight loss	two between, one within

RM-(M)ANOVA: Testing effects

Main goal in RM-(M)ANOVA: Test the significance of the effect of the within-factor on the DV (the treatment effect).

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k$$

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$$F = \frac{SS_B/df_B}{SS_W/df_W} = \frac{MS_B}{MS_W} \sim F(k-1, n-k)$$

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With repeated measures ANOVA we will now have two options (both using F tests):

- Run a univariate test.
- Run a multivariate test.

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ANOVA: $SS_T = SS_B + SS_W$ repeated measures ANOVA: $SS_T = SS_B + SS_{B_1} + SS_E$,

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where:

 $SS_T = \text{total SS}$

 SS_B = between treatments SS (the model SS)

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- SS_W = within treatments SS
- SS_{B_1} = between subjects SS (New! Consider subjects as a random factor!)

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where:

- SS_B = between treatments SS (the model SS)
- SS_W = within treatments SS
- SS_{B_1} = between subjects SS (New! Consider subjects as a random factor!)
- SS_E = residual SS

So, writing the SS in RM-ANOVA in the 'usual-ANOVA' way:



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Equivalently, and looking at the subjects blocking:

$$SS_{T} = SS_{\text{between subjects}} + SS_{\text{within subjects}}$$
$$= SS_{\text{between subjects}} + SS_{\text{between treatments}} + SS_{E}$$
The treatment effect is part of the within-subjects variance

Advantages:

 Reduction of within-group variance (error variance) by removing individual differences between subjects (by blocking on subjects).

greater power to test within-subjects effects

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Disadvantages:

- Order in which conditions are experienced may affect performance (practice effect, fatigue effect, carry-over effect).
- Model assumptions are harder to meet.
- More complicated model (of course...).

Goal: Study the effect of four drugs on reaction time. Each subject is tested using each drug.

		Drugs					
		1	2	3	4		
	1	30	28	16	34		
cts	2	14	18	10	22		
oje	3	24	20	18	30		
Sul	4	38	34	20	44		
	5	26	28	14	30		



Is there an effect of drug on reaction times?

You can compute each SS with the General Linear Model (Analyze>General Linear Model>Univariate)

Subject	Drug	Reaction
1	1	30
1	2	28
1	3	16
1	4	34
÷	÷	÷
5	1	26
5	2	28
5	3	14
5	4	30

Factors: Drug (fixed), Subject (random) **DV:** Reaction **Model:** Full factorial

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Tests of Between-Subjects Effects

_ Dependent Variable:reaction							
Source		Type III Sum of Squares	df	Mean Square	F	Siq.	
Intercept	Hypothesis	12400,200	1	12400,200	72,857	,001	
	Error	680,800	4	170,200 ^a			
drug	Hypothesis	SS _B 698,200	3	232,733	24,759	,000	
	Error	112,800	12	9,400 ^b			
subject	Hypothesis	SSB1 680,800	4	170,200	18,106	,000	
	Error	112,800	12	9,400 ^b			
drug * subject	Hypothesis	SSE 112,800	12	9,400			
	Error	,000	0	.°			

a. MS(subject)

b. MS(drug * subject)

c. MS(Error)

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drug * subject	Hypothesis	SS _E 112,800	12	9,400			
	Error	,000	0	.°			

a. MS(subject)

b. MS(drug * subject)

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Notes:

•
$$SS_W = SS_{B_1} + SS_E = 793.600$$

• $SS_{Model} = SS_B + SS_{B_1} = 1379.000$
• $SS_T = SS_{Model} + SS_E = 1491.800$
• Drug has a significant effect on reaction time $(F(3, 12) = 24.76, p < .001)$

Alternative way to do the univariate test: Analyze>General Linear Model>Repeated Measures.

Focus on the 'Sphericity Assumed' part (to be explained later)

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Analyze>General Linear Model>Repeated Measures.

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Source		Type III Sum of Squares	df	Mean Square	F	Sig.
Drug	Sphericity Assumed	698.200	3	232.733	24.759	.000
	Greenhouse-Geisser	698.200	1.815	384.763	24.759	.001
	Huynh-Feldt	698.200	3.000	232.733	24.759	.000
	Lower-bound	698.200	1.000	698.200	24.759	.008
Error(Drug)	Sphericity Assumed	112.800	12	9.400		
	Greenhouse-Geisser	112.800	7.258	15.540		
	Huynh-Feldt	112.800	12.000	9.400		
	Lower-bound	112.800	4.000	28.200		

Tests of Within-Subjects Effects

Measure:MEASURE_1

"RM-ANOVA removes individual differences between subjects from error SS_W "



Compare with the setting disregarding subjects blocking: •----



In general,

Source	SS	df	MS	F
Within-factor	SSB	k-1	SS_B/df_B	MS_B/MS_E
Between-factor	SS_{B_1}	n-1	SS_{B_1}/df_{B_1}	MS_{B_1}/MS_E
Error	SS_E	(k-1)(n-1)	SS_E/df_E	
Total	SS_T	nk-1		

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Within-factor	SS_B	k-1	SS_B/df_B	MS_B/MS_E
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Error	SS_E	(k-1)(n-1)	SS_E/df_E	
Total	SS_T	nk-1		

■ Subject SS (*SS*_{B1}): Variance between subjects.

In general,

Source	SS	df	MS	F
Within-factor	SSB	k-1	SS_B/df_B	MS_B/MS_E
Between-factor	SS_{B_1}	n-1	SS_{B_1}/df_{B_1}	MS_{B_1}/MS_E
Error	SS_E	(k-1)(n-1)	SS_E/df_E	
Total	SS_T	nk-1		

- Subject SS (*SS*_{B1}): Variance between subjects.
- Treatment SS (SS_B): Part of variance within subjects due to differences beetween treatments.
- Error SS (SS_E): Interaction part (to which extent subjects respond differently to treatments).

In general,

Source	SS	df	MS	F
Within-factor	SSB	k-1	SS_B/df_B	MS_B/MS_E
Between-factor	SS_{B_1}	n-1	SS_{B_1}/df_{B_1}	MS_{B_1}/MS_E
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The F tests in the table are the univariate tests for repeated measures analysis.

New approach: Regard each treatment level (or timepoint) as a DV.

In this way the repeated measures of one subject are a multivariate observation of the dependent variables.

For our	runnir	ng exar	nple		
Subject	Drug1	Drug2	Drug3	Drug4	
1	30	28	16	34	(multiveriate charged)
2	14	18	10	22 —	for subject 2
3	24	20	18	30	for subject 2
4	38	34	20	44	
5	26	28	14	30	

The main goal remains the same: Test

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k.$$

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Testing H_0 is similar to testing

$$H_0: \begin{pmatrix} \mu_1 - \mu_2 \\ \mu_2 - \mu_3 \\ \vdots \\ \mu_{k-1} - \mu_k \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}$$

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This test is multivariate because several (transformed) DVs are being compared in simultaneous. This is the generalization of the matched-pairs *t*-test

- There are several multivariate tests available; SPSS routinely reports four of them:
 - Wilk's Λ (use this test 'by default').
 - Pillai-Bartlett trace.
 - Hotelling-Lawley trace.
 - Roy's largest root.

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For our drug example:

Effect		Value	F	Hypothesis df	Error df	Sig.
Drug	Pillai's Trace	.977	28.412 ^a	3.000	2.000	.034
	Wilks' Lambda	.023	28.412ª	3.000	2.000	.034
	Hotelling's Trace	42.618	28.412 ^a	3.000	2.000	.034
	Roy's Largest Root	42.618	28.412 ^a	3.000	2.000	.034

Multivariate Tests^b

a. Exact statistic

b. Design: Intercept Within Subjects Design: Drug

Any of the tests rejects H_0 , i.e., drug has an effect on reaction times.

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RM-(M)ANOVA: Assumptions

Assumptions needed to perform the univariate/multivariate tests: For both tests:

- **Independence** of observations (between subjects).
- 2 Multivariate **normality** (i.e., each subgroup is normally distributed).
- Extra assumption for univariate *F*-test:
 - **3** Compound symmetry:
 - Variance is homogeneous across treatment levels (homogeneity of variances)

$$\sigma_{T_1}^2 = \sigma_{T_2}^2 = \dots = \sigma_{T_k}^2,$$

and

Covariances are homogeneous across treatment levels

$$\sigma_{T_1,T_2} = \sigma_{T_1,T_3} = \cdots = \sigma_{T_{k-1},T_k}.$$

RM-(M)ANOVA: Assumptions

- Compound symmetry is a highly restrictive condition which is not realistic in practice.
- It was shown that compound symmetry is sufficient but not necessary for the *F*-tests to be valid (i.e., compound symmetry is too strong).
- A more relaxed condition, under which the *F*-tests are valid, is sphericity:

All variances of all pairwise differences between pairs of repeated measures are equal:

$$\sigma_{T_1-T_2}^2 = \sigma_{T_1-T_3}^2 = \dots = \sigma_{T_{k-1}-T_k}^2.$$

SPSS tests the sphericity assumption with Mauchly's test. The goal is to NOT reject H_0 .

(This test is sensitive to departures from normality.)

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For our drug example:

Mauchly's Test of Sphericity^b

Measure:MEASURE_1

					Epsilon ^a			
Within Subjects Effect	Mauchly's W	Approx. Chi- Square	df	Sig.	Greenhouse- Geisser	Huynh-Feldt	Lower-bound	
Drug	.186	4.572	5	.495	.605	1.000	.333	
The null is not rejected, so sphericity holds								

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 - Introduce a correction in the univariate test, or
 - Use the multivariate test.

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- A: There are two possibilities:
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 - Use the multivariate test.

Q: How to correct lack of sphericity in the univariate test? A: Adjust the degrees of freedom of the *F* statistic before doing the tests (epsilon corrections).

In SPSS: Look at

- Greenhouse-Geisser correction.
- Huynh-Feldt correction.

Univariate test of $H_0: \mu_1 = \mu_2 = \cdots = \mu_k$ with epsilon corrections for the drug example:

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In general:

- Greenhouse-Geisser correction is too conservative:
 Overreject H₀.
- Huynh-Feldt correction too liberal: Underreject H₀.

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k$$

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A simplified rule which is half-way between Greenhouse-Geisser and Huynh-Feldt corrections is the following:

- 1 Compute the average of both epsilons (G-G and H-F), say $\overline{\varepsilon}$ 2 If...
 - $\bar{\varepsilon} > .7$: Use Huynh-Feldt correction.
 - $\bar{\varepsilon} < .7$ or nothing is known about sphericity: Use Greenhouse-Geisser correction.

Both methods are good for controlling for type I errors.

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'Sphericity' criterion

If sphericity is violated (ε < .7) and sample size is larger than k + 10: Use multivariate test statistics.
 (Because these statistics do not depend on this assumption, and they are

more powerful under these conditions.)

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■ If sphericity holds (ε > .7) or sample sizes are small: Use univariate test statistics

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'Missing values' criterion

Use univariate tests if there are missing values.

Both the univariate and multivariate tests are omnibus tests: They only detect overall effects.

We can understand better the within-factor effect by decomposing it into parts.

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Q: How to do it?

Both the univariate and multivariate tests are omnibus tests: They only detect overall effects.

We can understand better the within-factor effect by decomposing it into parts.

- Q: How to do it?
- A: Using contrasts.

SPSS has several contrasts available (Deviation, Helmert, Polynomial...).

Polynomial contrasts are interesting since it additively decomposes (and tests) the within-factor effect in linear, quadratic, cubic,..., effects: Trend analysis

('Additively' due to the orthogonality of polynomial contrasts.)

Tests of Within-Subjects Effects

Measure:MEASURE_1

Source		Type III Sum of Squares		df	Mean Square	F	Sig.
Drug	Sphericity Assumed		698.200	3	232.733	24.759	.000
	Greenhouse-Geisser		698.200	1.815	384.763	24.759	.001
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	Lower-bound		698.200	1.000	698.200	24.759	.008
Error(Drug)	Sphericity Assumed	7	112.800	12	9.400		
	Greenhouse-Geisser	/	112.800	7.258	15.540		
	Huynh-Feldt	/	112.800	12.000	9.400		
	Lower-bound		112.800	4.000	28.200		

Drug	g has a	a sig-
nifica	ant effe	ect.
But	what	else
can	we say	?)

Tests of Within-Subjects Contrasts

Measure:MEASURE 1

Source	Drug	T	ype III Sum of \$quares	df	Mean Square	F	Sia	
Drug	Linear		11.560	1	11.560	3.074	.154	
	Quadratic		/ 369.800	1	369.800	26.797	.007	
	Cubic		316.840	1	316.840	29.778	.005	
Error(Drug)	Linear		15.040	4	3.760			1
	Quadratic		55.200	4	13.800			L
	Cubic		42.560	4	10.640			

The effect of drug on reaction times is not significantly linear (but quadratic and cubic trends are sig.)

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Post hoc procedures

In RM-(M)ANOVA we can also use post hoc procedures.

These procedures allow doing multiple comparisons while preventing capitalization by chance (inflation of overall type I error).

In general:

- If sphericity is violated ($\varepsilon < .7$): Use Bonferroni procedure.
- If sphericity holds ($\varepsilon > .7$): Use Tukey procedure.

Note: Bonferroni is directly available in SPSS in 'Options' button of GLM's Repeated Measures menu; Tukey is not. Use of syntax files is needed (downloadable from SPSS' website, for example).

Post hoc procedures

For our drug example:

Pairwise Comparisons

Measure:MEASURE_1

					95% Confidence Interval for Difference ^a	
		Mean Difference (I-				
(I) Drug	(J) Drug	J)	Std. Error	Sig."	Lower Bound	Upper Bound
1	2	.800	1.625	1.000	-7.082	8.682
	3	10.800	2.577	.083	-1.700	23.300
	4	-5.600*	.748	.010	-9.230	-1.970
2	1	800	1.625	1.000	-8.682	7.082
	3	10.000	2.280	.071	-1.062	21.062
	4	-6.400	1.600	.097	-14.162	1.362
3	1	-10.800	2.577	.083	-23.300	1.700
	2	-10.000	2.280	.071	-21.062	1.062
	4	-16.400 [*]	2.227	.011	-27.204	-5.596
4	1	5.600*	.748	.010	1.970	9.230
	2	6.400	1.600	.097	-1.362	14.162
	3	16.400 [*]	2.227	.011	5.596	27.204

Based on estimated marginal means

a. Adjustment for multiple comparisons: Bonferroni.

*. The mean difference is significant at the .05 level.

Extending RM-ANOVA

- The RM-ANOVA situation covered here is the simplest: One within-factor (drug) only.
- The experimental design can be more complicated:
 - One between, one within design (Situation 2).
 - One between, two within design (Situation 3).
 - Two between, one within design (Situation 4).

...

RM-ANOVA can fit any of these extensions.